

Application of Random Matrix Theory to Future Wireless Networks

Romain Couillet^{1,2}

¹Alcatel-Lucent Chair on Flexible Radio, Supélec, Gif sur Yvette, FRANCE

²ST-Ericsson, Sophia-Antipolis, FRANCE

romain.couillet@supelec.fr

PhD defense

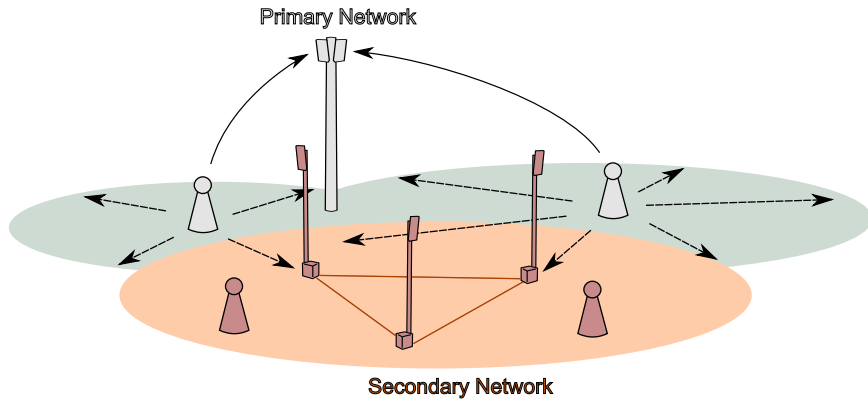
Supélec, November 12th, 2010.



- 1 Cognitive Radios
- 2 Exploration: Spectrum Hole Detection
- 3 Exploration: User Detection and Power Inference
- 4 Exploitation: Optimal Ergodic Rate
- 5 Perspectives and Conclusion

- 1 **Cognitive Radios**
- 2 Exploration: Spectrum Hole Detection
- 3 Exploration: User Detection and Power Inference
- 4 Exploitation: Optimal Ergodic Rate
- 5 Perspectives and Conclusion

General system model



Exploration and Exploitation

Cognitive Radio Setting:

- The secondary network is entitled to **reuse efficiently spectrum holes**,
 - so to **minimally interfere** the primary network;
 - so to **maximise the throughput** of secondary transmissions;
 - with little or no feedback from the primary network, i.e. **autonomously**.

Exploration and Exploitation

Cognitive Radio Setting:

- The secondary network is entitled to **reuse efficiently spectrum holes**,
 - so to **minimally interfere** the primary network;
 - so to **maximise the throughput** of secondary transmissions;
 - with little or no feedback from the primary network, i.e. **autonomously**.
- To this end, the secondary (or smart, cognitive) network must
 - **learn about the environment**: this is the **exploration phase**.
 - **communicate within the secondary network**: this is the **exploitation phase**.

Exploration and Exploitation

Cognitive Radio Setting:

- The secondary network is entitled to **reuse efficiently spectrum holes**,
 - so to **minimally interfere** the primary network;
 - so to **maximise the throughput** of secondary transmissions;
 - with little or no feedback from the primary network, i.e. **autonomously**.
- To this end, the secondary (or smart, cognitive) network must
 - **learn about the environment**: this is the **exploration phase**.
 - **communicate within the secondary network**: this is the **exploitation phase**.
- Both scenarios are inherently multi-dimensional:
 - efficient exploration requires large sensor networks;
 - networks may be composed of multiple users with possibly multiple antennas.

Exploration and Exploitation

Cognitive Radio Setting:

- The secondary network is entitled to **reuse efficiently spectrum holes**,
 - so to **minimally interfere** the primary network;
 - so to **maximise the throughput** of secondary transmissions;
 - with little or no feedback from the primary network, i.e. **autonomously**.
- To this end, the secondary (or smart, cognitive) network must
 - **learn about the environment**: this is the **exploration phase**.
 - **communicate within the secondary network**: this is the **exploitation phase**.
- Both scenarios are inherently multi-dimensional:
 - efficient exploration requires large sensor networks;
 - networks may be composed of multiple users with possibly multiple antennas.

We will discuss both exploration and exploitation phases as (possibly large) multivariate systems.

Outline

- 1 Cognitive Radios
- 2 Exploration: Spectrum Hole Detection**
- 3 Exploration: User Detection and Power Inference
- 4 Exploitation: Optimal Ergodic Rate
- 5 Perspectives and Conclusion

Formulation of the exploration problem

- We assume the scenario of a cognitive radio made of
 - a primary network of K transmitters, equipped with n_1, \dots, n_K antennas;
 - a secondary network in sensing mode, equipped of N collocated sensors.

Formulation of the exploration problem

- We assume the scenario of a cognitive radio made of
 - a primary network of K transmitters, equipped with n_1, \dots, n_K antennas;
 - a secondary network in sensing mode, equipped of N collocated sensors.
- At time m , the secondary network receives $\mathbf{y}^{(m)} \in \mathbb{C}^N$ as

$$\mathbf{y}^{(m)} = \sum_{k=1}^K \sqrt{P_k} \mathbf{H}_k \mathbf{x}_k^{(m)} + \sigma \mathbf{w}^{(m)}$$

Formulation of the exploration problem

- We assume the scenario of a cognitive radio made of
 - a primary network of K transmitters, equipped with n_1, \dots, n_K antennas;
 - a secondary network in sensing mode, equipped of N collocated sensors.
- At time m , the secondary network receives $\mathbf{y}^{(m)} \in \mathbb{C}^N$ as

$$\mathbf{y}^{(m)} = \sum_{k=1}^K \sqrt{P_k} \mathbf{H}_k \mathbf{x}_k^{(m)} + \sigma \mathbf{w}^{(m)}$$

- Depending on prior information, the secondary network will
 - try to infer the presence of a signal

Formulation of the exploration problem

- We assume the scenario of a cognitive radio made of
 - a primary network of K transmitters, equipped with n_1, \dots, n_K antennas;
 - a secondary network in sensing mode, equipped of N collocated sensors.
- At time m , the secondary network receives $\mathbf{y}^{(m)} \in \mathbb{C}^N$ as

$$\mathbf{y}^{(m)} = \sqrt{P}\mathbf{H}\mathbf{x}^{(m)} + \sigma\mathbf{w}^{(m)}$$

- Depending on prior information, the secondary network will
 - try to infer the presence of a signal (we will assume a unique transmitter)

Formulation of the exploration problem

- We assume the scenario of a cognitive radio made of
 - a primary network of K transmitters, equipped with n_1, \dots, n_K antennas;
 - a secondary network in sensing mode, equipped of N collocated sensors.
- At time m , the secondary network receives $\mathbf{y}^{(m)} \in \mathbb{C}^N$ as

$$\mathbf{y}^{(m)} = \sum_{k=1}^K \sqrt{P_k} \mathbf{H}_k \mathbf{x}_k^{(m)} + \sigma \mathbf{w}^{(m)}$$

- Depending on prior information, the secondary network will
 - try to infer the presence of a signal (we will assume a unique transmitter)
 - try to infer the transmit powers of the K transmitters.

Formulation of the exploration problem

- We assume the scenario of a cognitive radio made of
 - a primary network of K transmitters, equipped with n_1, \dots, n_K antennas;
 - a secondary network in sensing mode, equipped of N collocated sensors.
- At time m , the secondary network receives $\mathbf{y}^{(m)} \in \mathbb{C}^N$ as

$$\mathbf{y}^{(m)} = \sum_{k=1}^K \sqrt{P_k} \mathbf{H}_k \mathbf{x}_k^{(m)} + \sigma \mathbf{w}^{(m)}$$

- Depending on prior information, the secondary network will
 - try to infer the presence of a signal (we will assume a unique transmitter)
 - try to infer the transmit powers of the K transmitters.

This information allows for:

- the detection of **spectrum holes**;
- the evaluation of the **optimal secondary coverage**;
- *ideally*, the elaboration of a **'map' of space-frequency resources**.

Problem formulation

- We consider the model

$$\mathbf{y}^{(m)} = \begin{cases} \sigma \mathbf{w}^{(m)} & , (\mathcal{H}_0) \\ \sqrt{P} \mathbf{H} \mathbf{x}^{(m)} + \sigma \mathbf{w}^{(m)} & , (\mathcal{H}_1) \end{cases}$$

Problem formulation

- We consider the model

$$\mathbf{y}^{(m)} = \begin{cases} \sigma \mathbf{w}^{(m)} & , (\mathcal{H}_0) \\ \sqrt{P} \mathbf{H} \mathbf{x}^{(m)} + \sigma \mathbf{w}^{(m)} & , (\mathcal{H}_1) \end{cases}$$

- We wish to confront the hypotheses \mathcal{H}_0 and \mathcal{H}_1 given the data matrix $\mathbf{Y} \triangleq [\mathbf{y}^{(1)}, \dots, \mathbf{y}^{(M)}] \in \mathbb{C}^{N \times M}$.

Problem formulation

- We consider the model

$$\mathbf{y}^{(m)} = \begin{cases} \sigma \mathbf{w}^{(m)} & , (\mathcal{H}_0) \\ \sqrt{P} \mathbf{H} \mathbf{x}^{(m)} + \sigma \mathbf{w}^{(m)} & , (\mathcal{H}_1) \end{cases}$$

- We wish to confront the hypotheses \mathcal{H}_0 and \mathcal{H}_1 given the data matrix $\mathbf{Y} \triangleq [\mathbf{y}^{(1)}, \dots, \mathbf{y}^{(M)}] \in \mathbb{C}^{N \times M}$.
- We consider, in a **Bayesian framework**, the Neyman-Pearson test ratio

$$C(\mathbf{Y}) \triangleq \frac{P_{\mathcal{H}_1 | \mathbf{Y}, I}(\mathbf{Y})}{P_{\mathcal{H}_0 | \mathbf{Y}, I}(\mathbf{Y})}$$

with prior information I on $\mathbf{H}, \mathbf{x}^{(m)}, \sigma, \dots$

A Bayesian framework for cognitive radios

- We assume prior statistical and deterministic knowledge I on \mathbf{H}, σ, P

A Bayesian framework for cognitive radios

- We assume prior statistical and deterministic knowledge I on \mathbf{H}, σ, P
- Using the **maximum entropy principle** (MaxEnt), a prior $P_{(\mathbf{H}, \sigma, P)}(\mathbf{H}, \sigma, P)$ can be derived

$$P_{\mathbf{Y}|\mathcal{H}_i, I}(\mathbf{Y}) = \int_{(\mathbf{H}, \sigma, P)} P_{\mathbf{Y}|\mathcal{H}_i, I, \mathbf{H}, \sigma, P}(\mathbf{Y}) P_{(\mathbf{H}, \sigma, P)}(\mathbf{H}, \sigma, P) d(\mathbf{H}, \sigma, P)$$

A Bayesian framework for cognitive radios

- We assume prior statistical and deterministic knowledge I on \mathbf{H} , σ , P
- Using the **maximum entropy principle** (MaxEnt), a prior $P_{(\mathbf{H}, \sigma, P)}(\mathbf{H}, \sigma, P)$ can be derived

$$P_{\mathbf{Y}|\mathcal{H}_i, I}(\mathbf{Y}) = \int_{(\mathbf{H}, \sigma, P)} P_{\mathbf{Y}|\mathcal{H}_i, I, \mathbf{H}, \sigma, P}(\mathbf{Y}) P_{(\mathbf{H}, \sigma, P)}(\mathbf{H}, \sigma, P) d(\mathbf{H}, \sigma, P)$$

- In the following,
 - we derive the case $P = 1$, σ known and **the knowledge about \mathbf{H} conveys unitary invariance**
 - $E[\text{tr } \mathbf{H}\mathbf{H}^H]$ known: this is what we assume here;
 - $E[\mathbf{H}\mathbf{H}^H] = \mathbf{Q}$ unknown but such that $E[\text{tr } \mathbf{Q}]$ is known;
 - $\text{rank}(\mathbf{H}\mathbf{H}^H)$ known.
 - we compare alternative methods when $P = 1$ and σ are unknown.

Evaluation of $P_{\mathbf{Y}|\mathcal{H}_i, I}(\mathbf{Y})$

- by MaxEnt, \mathbf{X} , \mathbf{W} are standard Gaussian matrix with $X_{ij}, W_{ij} \sim \mathcal{CN}(0, 1)$.

Evaluation of $P_{\mathbf{Y}|\mathcal{H}_0, I}(\mathbf{Y})$

- by MaxEnt, \mathbf{X} , \mathbf{W} are standard Gaussian matrix with $X_{ij}, W_{ij} \sim \mathcal{CN}(0, 1)$.
- **Under \mathcal{H}_0 :**
 - $\mathbf{Y} = \sigma \mathbf{W}$

$$P_{\mathbf{Y}|\mathcal{H}_0, I}(\mathbf{Y}) = \frac{1}{(\pi\sigma^2)^{NM}} e^{-\frac{1}{\sigma^2} \text{tr} \mathbf{Y}\mathbf{Y}^H}.$$

Evaluation of $P_{\mathbf{Y}|\mathcal{H}_i,l}(\mathbf{Y})$

- by MaxEnt, \mathbf{X} , \mathbf{W} are standard Gaussian matrix with $X_{ij}, W_{ij} \sim \mathcal{CN}(0, 1)$.

- Under \mathcal{H}_0 :**

- $\mathbf{Y} = \sigma \mathbf{W}$

$$P_{\mathbf{Y}|\mathcal{H}_0,l}(\mathbf{Y}) = \frac{1}{(\pi\sigma^2)^{NM}} e^{-\frac{1}{\sigma^2} \text{tr} \mathbf{Y}\mathbf{Y}^H}.$$

- Under \mathcal{H}_1 :**

- $\mathbf{Y} = [\sqrt{P}\mathbf{H} \quad \sigma\mathbf{I}_N] \begin{bmatrix} \mathbf{X} \\ \mathbf{W} \end{bmatrix}$

$$P_{\mathbf{Y}|\mathcal{H}_1}(\mathbf{Y}) = \int_{\Sigma \geq 0} P_{\mathbf{Y}|\Sigma, \mathcal{H}_1}(\mathbf{Y}, \Sigma) P_{\Sigma}(\Sigma) d\Sigma$$

with $\Sigma = \mathbb{E}[\mathbf{y}^{(1)}\mathbf{y}^{(1)H}] = \mathbf{H}\mathbf{H}^H + \sigma^2\mathbf{I}_N$.

Evaluation of $P_{\mathbf{Y}|\mathcal{H}_i,l}(\mathbf{Y})$

- by MaxEnt, \mathbf{X} , \mathbf{W} are standard Gaussian matrix with $X_{ij}, W_{ij} \sim \mathcal{CN}(0, 1)$.

- Under \mathcal{H}_0 :**

- $\mathbf{Y} = \sigma \mathbf{W}$

$$P_{\mathbf{Y}|\mathcal{H}_0,l}(\mathbf{Y}) = \frac{1}{(\pi\sigma^2)^{NM}} e^{-\frac{1}{\sigma^2} \text{tr} \mathbf{Y}\mathbf{Y}^H}.$$

- Under \mathcal{H}_1 :**

- $\mathbf{Y} = [\sqrt{P}\mathbf{H} \quad \sigma\mathbf{I}_N] \begin{bmatrix} \mathbf{X} \\ \mathbf{W} \end{bmatrix}$

$$P_{\mathbf{Y}|\mathcal{H}_1}(\mathbf{Y}) = \int_{\Sigma \geq 0} P_{\mathbf{Y}|\Sigma, \mathcal{H}_1}(\mathbf{Y}, \Sigma) P_{\Sigma}(\Sigma) d\Sigma$$

with $\Sigma = \mathbb{E}[\mathbf{y}^{(1)}\mathbf{y}^{(1)H}] = \mathbf{H}\mathbf{H}^H + \sigma^2\mathbf{I}_N$.

From unitary invariance of \mathbf{H} , denoting $\Sigma = \mathbf{U}\mathbf{G}\mathbf{U}^H$, $\text{diag}(\mathbf{G}) = (g_1, \dots, g_n, \sigma^2, \dots, \sigma^2)$

$$P_{\mathbf{Y}|\mathcal{H}_1}(\mathbf{Y}) = \int_{\mathcal{U}(N) \times (\sigma^2, \infty)^n} P_{\mathbf{Y}|\mathbf{U}\mathbf{G}\mathbf{U}^H, \mathcal{H}_1}(\mathbf{Y}, \mathbf{U}, \mathbf{G}) P_{\mathbf{U}}(\mathbf{U}) P_{(g_1, \dots, g_n)}(g_1, \dots, g_n) d\mathbf{U} dg_1 \dots dg_n$$

where

Evaluation of $P_{\mathbf{Y}|\mathcal{H}_i,I}(\mathbf{Y})$

- by MaxEnt, \mathbf{X} , \mathbf{W} are standard Gaussian matrix with $X_{ij}, W_{ij} \sim \mathcal{CN}(0, 1)$.

- Under \mathcal{H}_0 :**

- $\mathbf{Y} = \sigma \mathbf{W}$

$$P_{\mathbf{Y}|\mathcal{H}_0,I}(\mathbf{Y}) = \frac{1}{(\pi\sigma^2)^{NM}} e^{-\frac{1}{\sigma^2} \text{tr} \mathbf{Y} \mathbf{Y}^H}.$$

- Under \mathcal{H}_1 :**

- $\mathbf{Y} = [\sqrt{P}\mathbf{H} \quad \sigma \mathbf{I}_N] \begin{bmatrix} \mathbf{X} \\ \mathbf{W} \end{bmatrix}$

$$P_{\mathbf{Y}|\mathcal{H}_1}(\mathbf{Y}) = \int_{\Sigma \geq 0} P_{\mathbf{Y}|\Sigma, \mathcal{H}_1}(\mathbf{Y}, \Sigma) P_{\Sigma}(\Sigma) d\Sigma$$

with $\Sigma = \mathbb{E}[\mathbf{y}^{(1)} \mathbf{y}^{(1)H}] = \mathbf{H} \mathbf{H}^H + \sigma^2 \mathbf{I}_N$.

From unitary invariance of \mathbf{H} , denoting $\Sigma = \mathbf{U} \mathbf{G} \mathbf{U}^H$, $\text{diag}(\mathbf{G}) = (g_1, \dots, g_n, \sigma^2, \dots, \sigma^2)$

$$P_{\mathbf{Y}|\mathcal{H}_1}(\mathbf{Y}) = \int_{\mathcal{U}(N) \times (\sigma^2, \infty)^n} P_{\mathbf{Y}|\mathbf{U} \mathbf{G} \mathbf{U}^H, \mathcal{H}_1}(\mathbf{Y}, \mathbf{U}, \mathbf{G}) P_{\mathbf{U}}(\mathbf{U}) P_{(g_1, \dots, g_n)}(g_1, \dots, g_n) d\mathbf{U} dg_1 \dots dg_n$$

where

- $P_{\mathbf{Y}|\mathbf{U} \mathbf{G} \mathbf{U}^H, \mathcal{H}_1}$ is Gaussian with zero mean and variance $\mathbf{U} \mathbf{G} \mathbf{U}^H$;

Evaluation of $P_{\mathbf{Y}|\mathcal{H}_i,l}(\mathbf{Y})$

- by MaxEnt, \mathbf{X} , \mathbf{W} are standard Gaussian matrix with $X_{ij}, W_{ij} \sim \mathcal{CN}(0, 1)$.

- Under \mathcal{H}_0 :**

- $\mathbf{Y} = \sigma \mathbf{W}$

$$P_{\mathbf{Y}|\mathcal{H}_0,l}(\mathbf{Y}) = \frac{1}{(\pi\sigma^2)^{NM}} e^{-\frac{1}{\sigma^2} \text{tr} \mathbf{Y}\mathbf{Y}^H}.$$

- Under \mathcal{H}_1 :**

- $\mathbf{Y} = [\sqrt{P}\mathbf{H} \quad \sigma\mathbf{I}_N] \begin{bmatrix} \mathbf{X} \\ \mathbf{W} \end{bmatrix}$

$$P_{\mathbf{Y}|\mathcal{H}_1}(\mathbf{Y}) = \int_{\Sigma \geq 0} P_{\mathbf{Y}|\Sigma, \mathcal{H}_1}(\mathbf{Y}, \Sigma) P_{\Sigma}(\Sigma) d\Sigma$$

with $\Sigma = \mathbb{E}[\mathbf{y}^{(1)}\mathbf{y}^{(1)H}] = \mathbf{H}\mathbf{H}^H + \sigma^2\mathbf{I}_N$.

From unitary invariance of \mathbf{H} , denoting $\Sigma = \mathbf{U}\mathbf{G}\mathbf{U}^H$, $\text{diag}(\mathbf{G}) = (g_1, \dots, g_n, \sigma^2, \dots, \sigma^2)$

$$P_{\mathbf{Y}|\mathcal{H}_1}(\mathbf{Y}) = \int_{\mathcal{U}(N) \times (\sigma^2, \infty)^n} P_{\mathbf{Y}|\mathbf{U}\mathbf{G}\mathbf{U}^H, \mathcal{H}_1}(\mathbf{Y}, \mathbf{U}, \mathbf{G}) P_{\mathbf{U}}(\mathbf{U}) P_{(g_1, \dots, g_n)}(g_1, \dots, g_n) d\mathbf{U} dg_1 \dots dg_n$$

where

- $P_{\mathbf{Y}|\mathbf{U}\mathbf{G}\mathbf{U}^H, \mathcal{H}_1}$ is Gaussian with zero mean and variance $\mathbf{U}\mathbf{G}\mathbf{U}^H$;
- $P_{\mathbf{U}}$ is a constant ($d\mathbf{U}$ is a Haar measure);

Evaluation of $P_{\mathbf{Y}|\mathcal{H}_i, I}(\mathbf{Y})$

- by MaxEnt, \mathbf{X} , \mathbf{W} are standard Gaussian matrix with $X_{ij}, W_{ij} \sim \mathcal{CN}(0, 1)$.

- Under \mathcal{H}_0 :**

- $\mathbf{Y} = \sigma \mathbf{W}$

$$P_{\mathbf{Y}|\mathcal{H}_0, I}(\mathbf{Y}) = \frac{1}{(\pi\sigma^2)^{NM}} e^{-\frac{1}{\sigma^2} \text{tr} \mathbf{Y} \mathbf{Y}^H}.$$

- Under \mathcal{H}_1 :**

- $\mathbf{Y} = [\sqrt{P} \mathbf{H} \quad \sigma \mathbf{I}_N] \begin{bmatrix} \mathbf{X} \\ \mathbf{W} \end{bmatrix}$

$$P_{\mathbf{Y}|\mathcal{H}_1}(\mathbf{Y}) = \int_{\Sigma \geq 0} P_{\mathbf{Y}|\Sigma, \mathcal{H}_1}(\mathbf{Y}, \Sigma) P_{\Sigma}(\Sigma) d\Sigma$$

with $\Sigma = \mathbb{E}[\mathbf{y}^{(1)} \mathbf{y}^{(1)H}] = \mathbf{H} \mathbf{H}^H + \sigma^2 \mathbf{I}_N$.

From unitary invariance of \mathbf{H} , denoting $\Sigma = \mathbf{U} \mathbf{G} \mathbf{U}^H$, $\text{diag}(\mathbf{G}) = (g_1, \dots, g_n, \sigma^2, \dots, \sigma^2)$

$$P_{\mathbf{Y}|\mathcal{H}_1}(\mathbf{Y}) = \int_{\mathcal{U}(N) \times (\sigma^2, \infty)^n} P_{\mathbf{Y}|\mathbf{U} \mathbf{G} \mathbf{U}^H, \mathcal{H}_1}(\mathbf{Y}, \mathbf{U}, \mathbf{G}) P_{\mathbf{U}}(\mathbf{U}) P_{(g_1, \dots, g_n)}(g_1, \dots, g_n) d\mathbf{U} dg_1 \dots dg_n$$

where

- $P_{\mathbf{Y}|\mathbf{U} \mathbf{G} \mathbf{U}^H, \mathcal{H}_1}$ is Gaussian with zero mean and variance $\mathbf{U} \mathbf{G} \mathbf{U}^H$;
- $P_{\mathbf{U}}$ is a constant ($d\mathbf{U}$ is a Haar measure);
- if \mathbf{H} is Gaussian, $P_{(g_1 - \sigma^2, \dots, g_n - \sigma^2)}$ is the joint eigenvalue distribution of a central Wishart;

Result in the Gaussian case, $n = 1$

R. Couillet, M. Debbah, "A Bayesian Framework for Collaborative Multi-Source Signal Sensing", IEEE Transactions on Signal Processing, vol. 58, no. 10, pp. 5186-5195, 2010.

Theorem (Neyman-Pearson test)

The ratio $C(\mathbf{Y})$ when the receiver knows $n = 1$, $P = 1$, $E[\frac{1}{N} \text{tr} \mathbf{H}\mathbf{H}^H] = 1$ and σ^2 , reads

$$C(\mathbf{Y}) = \frac{1}{N} \sum_{l=1}^N \frac{\sigma^{2(N+M-1)} e^{\sigma^2 + \frac{\lambda_l}{\sigma^2}}}{\prod_{\substack{i=1 \\ i \neq l}}^N (\lambda_l - \lambda_i)} J_{N-M-1}(\sigma^2, \lambda_l)$$

with $\lambda_1, \dots, \lambda_N$ the eigenvalues of $\mathbf{Y}\mathbf{Y}^H$ and where

$$J_k(x, y) \triangleq \int_x^{+\infty} t^k e^{-t - \frac{y}{t}} dt.$$

Result in the Gaussian case, $n = 1$

R. Couillet, M. Debbah, "A Bayesian Framework for Collaborative Multi-Source Signal Sensing", IEEE Transactions on Signal Processing, vol. 58, no. 10, pp. 5186-5195, 2010.

Theorem (Neyman-Pearson test)

The ratio $C(\mathbf{Y})$ when the receiver knows $n = 1$, $P = 1$, $E[\frac{1}{N} \text{tr} \mathbf{H}\mathbf{H}^H] = 1$ and σ^2 , reads

$$C(\mathbf{Y}) = \frac{1}{N} \sum_{l=1}^N \frac{\sigma^{2(N+M-1)} e^{\sigma^2 + \frac{\lambda_l}{\sigma^2}}}{\prod_{\substack{i=1 \\ i \neq l}}^N (\lambda_l - \lambda_i)} J_{N-M-1}(\sigma^2, \lambda_l)$$

with $\lambda_1, \dots, \lambda_N$ the eigenvalues of $\mathbf{Y}\mathbf{Y}^H$ and where

$$J_k(x, y) \triangleq \int_x^{+\infty} t^k e^{-t - \frac{y}{t}} dt.$$

- non trivial dependency on $\lambda_1, \dots, \lambda_N$
- contrary to energy detector, $\sum_i \lambda_i$ is not a sufficient statistic;

Result in the Gaussian case, $n = 1$

R. Couillet, M. Debbah, "A Bayesian Framework for Collaborative Multi-Source Signal Sensing", IEEE Transactions on Signal Processing, vol. 58, no. 10, pp. 5186-5195, 2010.

Theorem (Neyman-Pearson test)

The ratio $C(\mathbf{Y})$ when the receiver knows $n = 1$, $P = 1$, $E[\frac{1}{N} \text{tr} \mathbf{H}\mathbf{H}^H] = 1$ and σ^2 , reads

$$C(\mathbf{Y}) = \frac{1}{N} \sum_{l=1}^N \frac{\sigma^{2(N+M-1)} e^{\sigma^2 + \frac{\lambda_l}{\sigma^2}}}{\prod_{\substack{i=1 \\ i \neq l}}^N (\lambda_l - \lambda_i)} J_{N-M-1}(\sigma^2, \lambda_l)$$

with $\lambda_1, \dots, \lambda_N$ the eigenvalues of $\mathbf{Y}\mathbf{Y}^H$ and where

$$J_k(x, y) \triangleq \int_x^{+\infty} t^k e^{-t - \frac{y}{t}} dt.$$

- non trivial dependency on $\lambda_1, \dots, \lambda_N$
- contrary to energy detector, $\sum_i \lambda_i$ is not a sufficient statistic;
- **integration over σ^2** (or P when $P \neq 1$) is difficult.

Comparison to energy detector

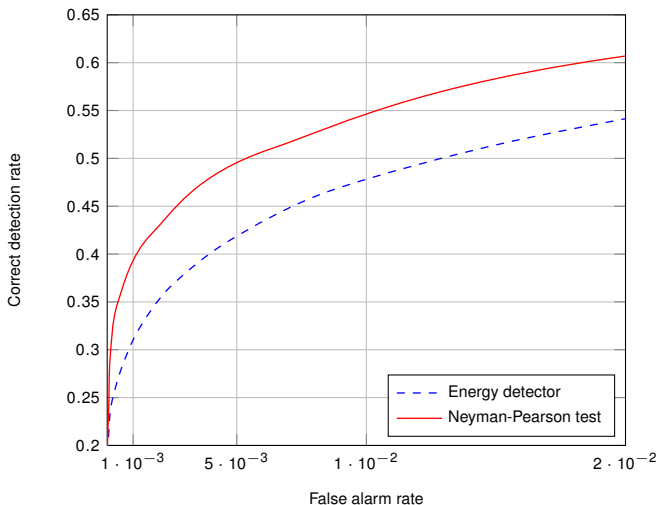


Figure: ROC curve for single-source detection, $K = 1$, $N = 4$, $M = 8$, $\text{SNR} = -3$ dB, FAR range of practical interest, with signal power $E = 0$ dBm, either known or unknown at the receiver.

Comparison to energy detector

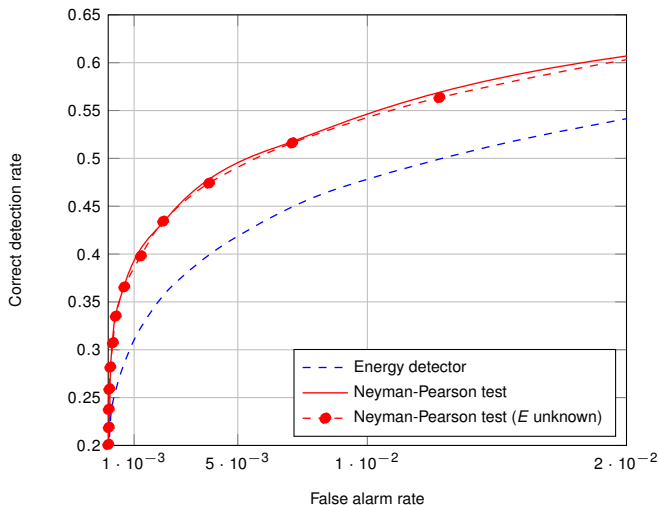


Figure: ROC curve for single-source detection, $K = 1$, $N = 4$, $M = 8$, $\text{SNR} = -3$ dB, FAR range of practical interest, with signal power $E = 0$ dBm, either known or unknown at the receiver.

Unknown power and noise variances

- Bayesian approaches:

$$P_{\mathbf{Y}|\mathcal{H}_i, I}(\mathbf{Y}) = \int_{\mathbb{R}_+^2} P_{\mathbf{Y}|\mathcal{H}_i, \sigma, P}(\mathbf{Y}) P_{(\sigma, P)}(\sigma, P) d(\sigma, P)$$

- limited by **computational complexity** (two-dimension numerical integration);
- inconsistence in MaxEnt **uninformative priors** on σ, P .

Unknown power and noise variances

- Bayesian approaches:

$$P_{\mathbf{Y}|\mathcal{H}_i, I}(\mathbf{Y}) = \int_{\mathbb{R}_+^2} P_{\mathbf{Y}|\mathcal{H}_i, \sigma, P}(\mathbf{Y}) P_{(\sigma, P)}(\sigma, P) d(\sigma, P)$$

- limited by **computational complexity** (two-dimension numerical integration);
- inconsistence in MaxEnt **uninformative priors** on σ, P .
- non-parametric GLRT approach:

$$C_{\text{GLRT}}(\mathbf{Y}) \triangleq \frac{\sup_{\mathbf{H}, \sigma} P_{\mathbf{Y}|\mathcal{H}_1}(\mathbf{Y})}{\sup_{\sigma} P_{\mathbf{Y}|\mathcal{H}_0}(\mathbf{Y})}$$

- $C_{\text{GLRT}}(\mathbf{Y})$ expresses as a monotonic function of $\frac{\max_j \lambda_j}{\frac{1}{N} \sum_i \lambda_i}$;
- excludes prior information** on \mathbf{H}, σ, P .

Unknown power and noise variances

- Bayesian approaches:

$$P_{\mathbf{Y}|\mathcal{H}_i, I}(\mathbf{Y}) = \int_{\mathbb{R}_+^2} P_{\mathbf{Y}|\mathcal{H}_i, \sigma, P}(\mathbf{Y}) P_{(\sigma, P)}(\sigma, P) d(\sigma, P)$$

- limited by **computational complexity** (two-dimension numerical integration);
- inconsistence in MaxEnt **uninformative priors** on σ, P .
- non-parametric GLRT approach:

$$C_{\text{GLRT}}(\mathbf{Y}) \triangleq \frac{\sup_{\mathbf{H}, \sigma} P_{\mathbf{Y}|\mathcal{H}_1}(\mathbf{Y})}{\sup_{\sigma} P_{\mathbf{Y}|\mathcal{H}_0}(\mathbf{Y})}$$

- $C_{\text{GLRT}}(\mathbf{Y})$ expresses as a monotonic function of $\frac{\max_i \lambda_i}{\frac{1}{N} \sum_i \lambda_i}$;
- excludes prior information** on \mathbf{H}, σ, P .
- ad-hoc methods, such as conditioning number:

$$C_{\text{cond}}(\mathbf{Y}) \triangleq \frac{\max_i \lambda_i}{\min_i \lambda_i}$$

- based on large random matrix considerations: under \mathcal{H}_0 , as $N/M \rightarrow c$

$$\frac{\max_i \lambda_i}{\min_i \lambda_i} \xrightarrow{\text{a.s.}} \frac{(1 + \sqrt{c})^2}{(1 - \sqrt{c})^2}$$

- totally **empirical**.

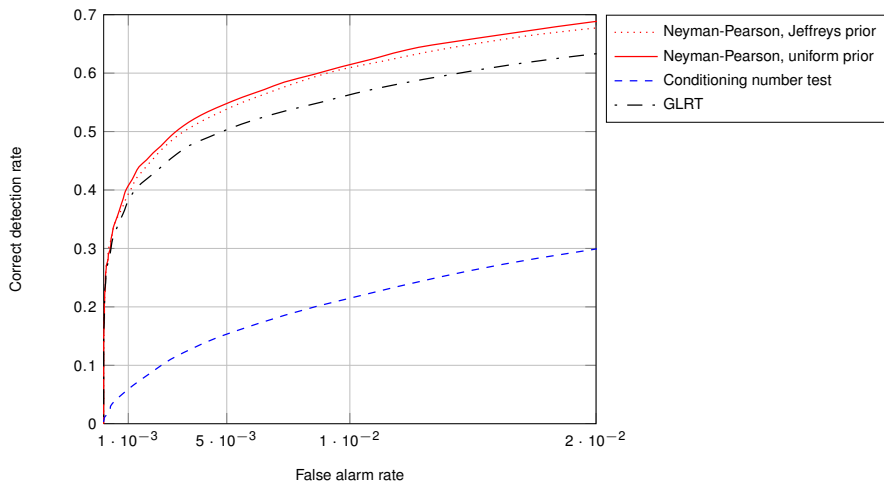
Performance comparison for unknown σ^2, P 

Figure: ROC curve for *a priori* unknown σ^2 of the Neyman-Pearson test, conditioning number method and GLRT, $K = 1$, $N = 4$, $M = 8$, SNR = 0 dB. For the Neyman-Pearson test, both uniform and Jeffreys prior, with exponent $\beta = 1$, are provided.

Outline

- 1 Cognitive Radios
- 2 Exploration: Spectrum Hole Detection
- 3 Exploration: User Detection and Power Inference**
- 4 Exploitation: Optimal Ergodic Rate
- 5 Perspectives and Conclusion

Problem Statement

- We now consider the model

$$\mathbf{y}^{(m)} = \sum_{k=1}^K \sqrt{P_k} \mathbf{H}_k \mathbf{x}_k^{(m)} + \sigma \mathbf{w}^{(m)}$$

and wish to infer P_1, \dots, P_K .

Problem Statement

- We now consider the model

$$\mathbf{y}^{(m)} = \sum_{k=1}^K \sqrt{P_k} \mathbf{H}_k \mathbf{x}_k^{(m)} + \sigma \mathbf{w}^{(m)}$$

and wish to *infer* P_1, \dots, P_K .

- With $\mathbf{Y} = [\mathbf{y}^{(1)}, \dots, \mathbf{y}^{(M)}]$, this can be rewritten

$$\mathbf{Y} = \sum_{k=1}^K \sqrt{P_k} \mathbf{H}_k \mathbf{X}_k + \sigma \mathbf{W}$$

Problem Statement

- We now consider the model

$$\mathbf{y}^{(m)} = \sum_{k=1}^K \sqrt{P_k} \mathbf{H}_k \mathbf{x}_k^{(m)} + \sigma \mathbf{w}^{(m)}$$

and wish to *infer* P_1, \dots, P_K .

- With $\mathbf{Y} = [\mathbf{y}^{(1)}, \dots, \mathbf{y}^{(M)}]$, this can be rewritten

$$\mathbf{Y} = \sum_{k=1}^K \sqrt{P_k} \mathbf{H}_k \mathbf{X}_k + \sigma \mathbf{W} = \underbrace{\left[\sqrt{P_1} \mathbf{H}_1 \quad \dots \quad \sqrt{P_K} \mathbf{H}_K \right]}_{\triangleq \mathbf{H} \mathbf{P}^{\frac{1}{2}}} \underbrace{\begin{bmatrix} \mathbf{X}_1 \\ \vdots \\ \mathbf{X}_K \end{bmatrix}}_{\triangleq \mathbf{X}} + \sigma \mathbf{W}$$

Problem Statement

- We now consider the model

$$\mathbf{y}^{(m)} = \sum_{k=1}^K \sqrt{P_k} \mathbf{H}_k \mathbf{x}_k^{(m)} + \sigma \mathbf{w}^{(m)}$$

and wish to **infer** P_1, \dots, P_K .

- With $\mathbf{Y} = [\mathbf{y}^{(1)}, \dots, \mathbf{y}^{(M)}]$, this can be rewritten

$$\mathbf{Y} = \sum_{k=1}^K \sqrt{P_k} \mathbf{H}_k \mathbf{X}_k + \sigma \mathbf{W} = \underbrace{\begin{bmatrix} \sqrt{P_1} \mathbf{H}_1 & \dots & \sqrt{P_K} \mathbf{H}_K \end{bmatrix}}_{\triangleq \mathbf{H} \mathbf{P}^{\frac{1}{2}}} \underbrace{\begin{bmatrix} \mathbf{X}_1 \\ \vdots \\ \mathbf{X}_K \end{bmatrix}}_{\triangleq \mathbf{X}} + \sigma \mathbf{W} = \begin{bmatrix} \mathbf{H} \mathbf{P}^{\frac{1}{2}} & \sigma \mathbf{I}_N \end{bmatrix} \begin{bmatrix} \mathbf{X} \\ \mathbf{W} \end{bmatrix}.$$

Problem Statement

- We now consider the model

$$\mathbf{y}^{(m)} = \sum_{k=1}^K \sqrt{P_k} \mathbf{H}_k \mathbf{x}_k^{(m)} + \sigma \mathbf{w}^{(m)}$$

and wish to infer P_1, \dots, P_K .

- With $\mathbf{Y} = [\mathbf{y}^{(1)}, \dots, \mathbf{y}^{(M)}]$, this can be rewritten

$$\mathbf{Y} = \sum_{k=1}^K \sqrt{P_k} \mathbf{H}_k \mathbf{X}_k + \sigma \mathbf{W} = \underbrace{\begin{bmatrix} \sqrt{P_1} \mathbf{H}_1 & \dots & \sqrt{P_K} \mathbf{H}_K \end{bmatrix}}_{\triangleq \mathbf{H} \mathbf{P}^{\frac{1}{2}}} \underbrace{\begin{bmatrix} \mathbf{X}_1 \\ \vdots \\ \mathbf{X}_K \end{bmatrix}}_{\triangleq \mathbf{X}} + \sigma \mathbf{W} = \begin{bmatrix} \mathbf{H} \mathbf{P}^{\frac{1}{2}} & \sigma \mathbf{I}_N \end{bmatrix} \begin{bmatrix} \mathbf{X} \\ \mathbf{W} \end{bmatrix}.$$

- If \mathbf{H} , $(\mathbf{X}^T \mathbf{W}^T)$ are unitarily invariant, \mathbf{Y} is unitarily invariant.

Problem Statement

- We now consider the model

$$\mathbf{y}^{(m)} = \sum_{k=1}^K \sqrt{P_k} \mathbf{H}_k \mathbf{x}_k^{(m)} + \sigma \mathbf{w}^{(m)}$$

and wish to infer P_1, \dots, P_K .

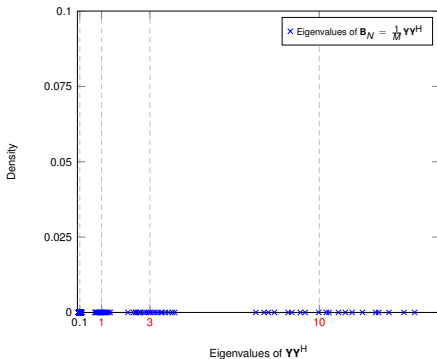
- With $\mathbf{Y} = [\mathbf{y}^{(1)}, \dots, \mathbf{y}^{(M)}]$, this can be rewritten

$$\mathbf{Y} = \sum_{k=1}^K \sqrt{P_k} \mathbf{H}_k \mathbf{X}_k + \sigma \mathbf{W} = \underbrace{\begin{bmatrix} \sqrt{P_1} \mathbf{H}_1 & \dots & \sqrt{P_K} \mathbf{H}_K \end{bmatrix}}_{\triangleq \mathbf{H} \mathbf{P}^{\frac{1}{2}}} \underbrace{\begin{bmatrix} \mathbf{X}_1 \\ \vdots \\ \mathbf{X}_K \end{bmatrix}}_{\triangleq \mathbf{X}} + \sigma \mathbf{W} = \begin{bmatrix} \mathbf{H} \mathbf{P}^{\frac{1}{2}} & \sigma \mathbf{I}_N \end{bmatrix} \begin{bmatrix} \mathbf{X} \\ \mathbf{W} \end{bmatrix}.$$

- If \mathbf{H} , $(\mathbf{X}^T \mathbf{W}^T)$ are unitarily invariant, \mathbf{Y} is unitarily invariant.

Most information about P_1, \dots, P_K is contained in the eigenvalues of $\mathbf{B}_N \triangleq \frac{1}{M} \mathbf{Y} \mathbf{Y}^H$.

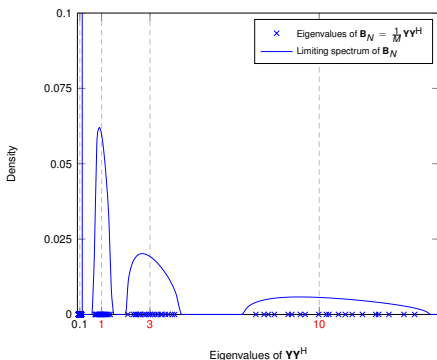
From small to large system analysis



The classical approach requires to evaluate $P_{P_1, \dots, P_K | Y}$

- assuming Gaussian parameters, this is **similar to previous calculus**
- leads to a **sum of two-dimensional integrals**
- prohibitively expensive to evaluate even for small N , n_k , M

From small to large system analysis



Assuming dimensions N , n_K , M grow large, **large dimensional random matrix theory** provides

- a link between:
 - **the “observation”**: the limiting spectral distribution (l.s.d.) of \mathbf{B}_N ;
 - **the “hidden parameters”**: the powers P_1, \dots, P_K , i.e. the l.s.d. of \mathbf{P} .
- **consistent estimators** of the hidden parameters.

Limiting spectrum of the sample covariance matrix

- **Definition.** The Stieltjes transform $m_F(z)$, of a distribution function F is defined as

$$m_F(z) = \int \frac{1}{\omega - z} dF(\omega).$$

Knowing the Stieltjes transform of F is equivalent to knowing F (similar to Fourier transform).

Limiting spectrum of the sample covariance matrix

- **Definition.** The Stieltjes transform $m_F(z)$, of a distribution function F is defined as

$$m_F(z) = \int \frac{1}{\omega - z} dF(\omega).$$

Knowing the Stieltjes transform of F is equivalent to knowing F (similar to Fourier transform).

- for simplicity, consider the *sample covariance matrix* model

$$\mathbf{Y} \triangleq \mathbf{T}^{\Delta} \frac{1}{\sqrt{2}} \mathbf{X} \in \mathbb{C}^{N \times n}, \quad \mathbf{B}_N = \frac{1}{n} \mathbf{Y} \mathbf{Y}^H \in \mathbb{C}^{N \times N}, \quad \mathbf{B}_n = \frac{1}{n} \mathbf{Y}^H \mathbf{Y} \in \mathbb{C}^{n \times n}$$

where $\mathbf{T} \in \mathbb{C}^{N \times N}$ has eigenvalues t_1, \dots, t_K , t_k with multiplicity N_k and $\mathbf{X} \in \mathbb{C}^{N \times n}$ is i.i.d. zero mean, variance 1.

Limiting spectrum of the sample covariance matrix

- **Definition.** The Stieltjes transform $m_F(z)$, of a distribution function F is defined as

$$m_F(z) = \int \frac{1}{\omega - z} dF(\omega).$$

Knowing the Stieltjes transform of F is equivalent to knowing F (similar to Fourier transform).

- for simplicity, consider the *sample covariance matrix* model

$$\mathbf{Y} \triangleq \mathbf{T}^{\frac{1}{2}} \mathbf{X} \in \mathbb{C}^{N \times n}, \quad \mathbf{B}_N = \frac{1}{n} \mathbf{Y} \mathbf{Y}^H \in \mathbb{C}^{N \times N}, \quad \underline{\mathbf{B}}_N = \frac{1}{n} \mathbf{Y}^H \mathbf{Y} \in \mathbb{C}^{n \times n}$$

where $\mathbf{T} \in \mathbb{C}^{N \times N}$ has eigenvalues t_1, \dots, t_K , t_k with multiplicity N_k and $\mathbf{X} \in \mathbb{C}^{N \times n}$ is i.i.d. zero mean, variance 1.

- If $F^{\mathbf{T}} \Rightarrow T$, then $m_{\mathbf{B}_N}(z) = m_{\underline{\mathbf{B}}_N}(z) \xrightarrow{\text{a.s.}} m_F(z)$ such that

$$m_T(-1/m_{\underline{F}}(z)) = -z m_{\underline{F}}(z) m_F(z)$$

with $m_{\underline{F}}(z) = c m_F(z) + (c - 1) \frac{1}{z}$ and $N/n \rightarrow c$.

Complex integration

- From Cauchy integral formula, denoting \mathcal{C}_k a contour enclosing **only** t_k ,

$$t_k = \frac{1}{2\pi i} \oint_{\mathcal{C}_k} \frac{\omega}{\omega - t_k} d\omega$$

Complex integration

- From Cauchy integral formula, denoting C_k a contour enclosing **only** t_k ,

$$t_k = \frac{1}{2\pi i} \oint_{C_k} \frac{\omega}{\omega - t_k} d\omega = \frac{1}{2\pi i} \oint_{C_k} \frac{1}{N_k} \sum_{j=1}^K N_j \frac{\omega}{\omega - t_j} d\omega$$

Complex integration

- From Cauchy integral formula, denoting C_k a contour enclosing **only** t_k ,

$$t_k = \frac{1}{2\pi i} \oint_{C_k} \frac{\omega}{\omega - t_k} d\omega = \frac{1}{2\pi i} \oint_{C_k} \frac{1}{N_k} \sum_{j=1}^K N_j \frac{\omega}{\omega - t_j} d\omega = \frac{N}{2\pi i N_k} \oint_{C_k} \omega m_T(\omega) d\omega.$$

Complex integration

- From Cauchy integral formula, denoting C_k a contour enclosing **only** t_k ,

$$t_k = \frac{1}{2\pi i} \oint_{C_k} \frac{\omega}{\omega - t_k} d\omega = \frac{1}{2\pi i} \oint_{C_k} \frac{1}{N_k} \sum_{j=1}^K N_j \frac{\omega}{\omega - t_j} d\omega = \frac{N}{2\pi i N_k} \oint_{C_k} \omega m_T(\omega) d\omega.$$

- After the variable change $\omega = -1/m_F(z)$,

$$t_k = \frac{N}{N_k} \frac{1}{2\pi i} \oint_{C_{E,k}} z m_F(z) \frac{m'_F(z)}{m_F^2(z)} dz,$$

Complex integration

- From Cauchy integral formula, denoting \mathcal{C}_k a contour enclosing **only** t_k ,

$$t_k = \frac{1}{2\pi i} \oint_{\mathcal{C}_k} \frac{\omega}{\omega - t_k} d\omega = \frac{1}{2\pi i} \oint_{\mathcal{C}_k} \frac{1}{N_k} \sum_{j=1}^K N_j \frac{\omega}{\omega - t_j} d\omega = \frac{N}{2\pi i N_k} \oint_{\mathcal{C}_k} \omega m_T(\omega) d\omega.$$

- After the variable change $\omega = -1/m_F(z)$,

$$t_k = \frac{N}{N_k} \frac{1}{2\pi i} \oint_{\mathcal{C}_{E,k}} z m_F(z) \frac{m'_F(z)}{m_F^2(z)} dz,$$

- When the system dimensions are large,

$$m_F(z) \simeq m_{\mathbf{B}_N}(z) \triangleq \frac{1}{N} \sum_{k=1}^N \frac{1}{\lambda_k - z}, \quad \text{with } (\lambda_1, \dots, \lambda_N) = \text{eig}(\mathbf{B}_N) = \text{eig}(\mathbf{Y}\mathbf{Y}^H).$$

Complex integration

- From Cauchy integral formula, denoting \mathcal{C}_k a contour enclosing **only** t_k ,

$$t_k = \frac{1}{2\pi i} \oint_{\mathcal{C}_k} \frac{\omega}{\omega - t_k} d\omega = \frac{1}{2\pi i} \oint_{\mathcal{C}_k} \frac{1}{N_k} \sum_{j=1}^K N_j \frac{\omega}{\omega - t_j} d\omega = \frac{N}{2\pi i N_k} \oint_{\mathcal{C}_k} \omega m_T(\omega) d\omega.$$

- After the variable change $\omega = -1/m_F(z)$,

$$t_k = \frac{N}{N_k} \frac{1}{2\pi i} \oint_{\mathcal{C}_{E,k}} z m_F(z) \frac{m'_F(z)}{m_F^2(z)} dz,$$

- When the system dimensions are large,

$$m_F(z) \simeq m_{\mathbf{B}_N}(z) \triangleq \frac{1}{N} \sum_{k=1}^N \frac{1}{\lambda_k - z}, \quad \text{with } (\lambda_1, \dots, \lambda_N) = \text{eig}(\mathbf{B}_N) = \text{eig}(\mathbf{Y}\mathbf{Y}^H).$$

- Dominated convergence arguments then show

$$t_k - \hat{t}_k \xrightarrow{\text{a.s.}} 0 \quad \text{with} \quad \hat{t}_k = \frac{N}{N_k} \frac{1}{2\pi i} \oint_{\mathcal{C}_{E,k}} z m_{\mathbf{B}_N}(z) \frac{m'_{\mathbf{B}_N}(z)}{m_{\mathbf{B}_N}^2(z)} dz$$

Complex integration

- From Cauchy integral formula, denoting \mathcal{C}_k a contour enclosing **only** t_k ,

$$t_k = \frac{1}{2\pi i} \oint_{\mathcal{C}_k} \frac{\omega}{\omega - t_k} d\omega = \frac{1}{2\pi i} \oint_{\mathcal{C}_k} \frac{1}{N_k} \sum_{j=1}^K N_j \frac{\omega}{\omega - t_j} d\omega = \frac{N}{2\pi i N_k} \oint_{\mathcal{C}_k} \omega m_T(\omega) d\omega.$$

- After the variable change $\omega = -1/m_F(z)$,

$$t_k = \frac{N}{N_k} \frac{1}{2\pi i} \oint_{\mathcal{C}_{E,k}} z m_F(z) \frac{m'_F(z)}{m_F^2(z)} dz,$$

- When the system dimensions are large,

$$m_F(z) \simeq m_{\mathbf{B}_N}(z) \triangleq \frac{1}{N} \sum_{k=1}^N \frac{1}{\lambda_k - z}, \quad \text{with } (\lambda_1, \dots, \lambda_N) = \text{eig}(\mathbf{B}_N) = \text{eig}(\mathbf{Y}\mathbf{Y}^H).$$

- Dominated convergence arguments then show

$$t_k - \hat{t}_k \xrightarrow{\text{a.s.}} 0 \quad \text{with} \quad \hat{t}_k = \frac{N}{N_k} \frac{1}{2\pi i} \oint_{\mathcal{C}_{E,k}} z m_{\mathbf{B}_N}(z) \frac{m'_{\mathbf{B}_N}(z)}{m_{\mathbf{B}_N}^2(z)} dz = \frac{n}{N_k} \sum_{m \in \mathcal{N}_k} (\lambda_m - \mu_m)$$

with \mathcal{N}_k the indexes of cluster k and $\mu_1 < \dots < \mu_N$ are the ordered eigenvalues of the matrix $\text{diag}(\boldsymbol{\lambda}) - \frac{1}{n} \sqrt{\boldsymbol{\lambda}} \sqrt{\boldsymbol{\lambda}}^T$, $\boldsymbol{\lambda} = (\lambda_1, \dots, \lambda_N)^T$.

Application to the current model

R. Couillet, J. W. Silverstein, Z. Bai, M. Debbah, "Eigen-Inference for Energy Estimation of Multiple Sources," *to appear in IEEE Transactions on Information Theory*, 2010.

- Extending \mathbf{Y} with zeros, our model is a "double sample covariance matrix"

$$\underbrace{\mathbf{Y}}_{(N+n) \times M} = \underbrace{\begin{bmatrix} \mathbf{H}\mathbf{P}^1 & \sigma\mathbf{I}_N \\ \mathbf{0} & \mathbf{0} \end{bmatrix}}_{(N+n) \times (N+n)} \underbrace{\begin{bmatrix} \mathbf{X} \\ \mathbf{W} \end{bmatrix}}_{(N+n) \times M}.$$

- Limiting distribution of $\frac{1}{M} \mathbf{Y}\mathbf{Y}^H$

Theorem (l.s.d. of \mathbf{B}_N)

Let $\mathbf{B}_N = \frac{1}{M} \mathbf{Y}\mathbf{Y}^H$ with eigenvalues $\lambda_1, \dots, \lambda_N$. Denote $m_{\mathbf{B}_N}(z) \triangleq \frac{1}{M} \sum_{k=1}^M \frac{1}{\lambda_k - z}$, with $\lambda_i = 0$ for $i > N$. Then, for $M/N \rightarrow c$, $N/n_k \rightarrow c_k$, $N/n \rightarrow c_0$, for any $z \in \mathbb{C}^+$,

$$m_{\mathbf{B}_N}(z) \xrightarrow{\text{a.s.}} m_{\underline{F}}(z)$$

with $m_{\underline{F}}(z)$ the unique solution in \mathbb{C}^+ of

$$\frac{1}{m_{\underline{F}}(z)} = -\sigma^2 + \frac{1}{f(z)} \left[\frac{c_0 - 1}{c_0} + m_P \left(-\frac{1}{f(z)} \right) \right], \quad \text{with } f(z) = (c - 1)m_{\underline{F}}(z) - czm_{\underline{F}}(z)^2.$$

Application to the current model (2)

R. Couillet, J. W. Silverstein, Z. Bai, M. Debbah, "Eigen-Inference for Energy Estimation of Multiple Sources," to appear in IEEE Trans. on Inf. Theory, 2010.

- estimator calculus

Theorem (Estimator of P_1, \dots, P_K)

Let $\mathbf{B}_N \in \mathbb{C}^{N \times N}$ be defined as in Theorem 2, and $\boldsymbol{\lambda} = (\lambda_1, \dots, \lambda_N)$, $\lambda_1 < \dots < \lambda_N$. Assume that asymptotic **cluster separability condition** is fulfilled for some k . Then, as $N, n, M \rightarrow \infty$,

$$\hat{P}_k - P_k \xrightarrow{\text{a.s.}} 0,$$

where

$$\hat{P}_k = \frac{NM}{n_k(M-N)} \sum_{i \in \mathcal{N}_k} (\eta_i - \mu_i)$$

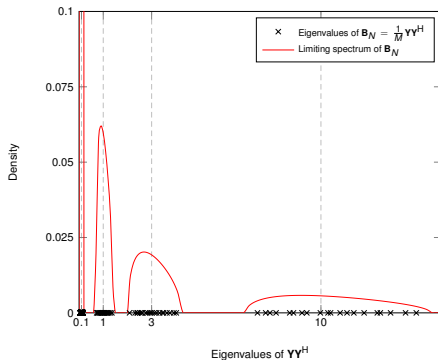
with \mathcal{N}_k the set indexing the eigenvalues in cluster k of F , $\eta_1 < \dots < \eta_N$ the eigenvalues of $\text{diag}(\boldsymbol{\lambda}) - \frac{1}{N} \sqrt{\boldsymbol{\lambda}} \sqrt{\boldsymbol{\lambda}}^T$ and $\mu_1 < \dots < \mu_N$ the eigenvalues of $\text{diag}(\boldsymbol{\lambda}) - \frac{1}{M} \sqrt{\boldsymbol{\lambda}} \sqrt{\boldsymbol{\lambda}}^T$.

Remarks

- solution is computationally simple, **explicit**, and the final formula compact.

Remarks

- solution is computationally simple, **explicit**, and the final formula compact.
- cluster separability condition is fundamental. This requires
 - for all other parameters fixed, the P_k cannot be too close to one another: **source separation problem**.
 - for all other parameters fixed, σ^2 must be kept low: **low SNR undecidability problem**.
 - for all other parameters fixed, M/N cannot be too low: **sample deficiency issue** (not such an issue though).
 - for all other parameters fixed, N/n cannot be too low: **diversity issue**.
- **exact spectrum separability** is an essential ingredient (known for very few models to this day).



Simulations

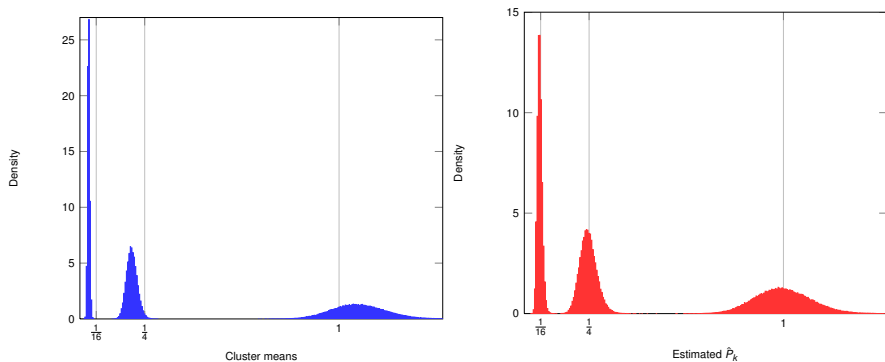


Figure: Histogram of the cluster-mean approach and of \hat{P}_k for $k \in \{1, 2, 3\}$, $P_1 = 1/16$, $P_2 = 1/4$, $P_3 = 1$, $n_1 = n_2 = n_3 = 4$ antennas per user, $N = 24$ sensors, $M = 128$ samples and SNR = 20 dB.

Performance comparison

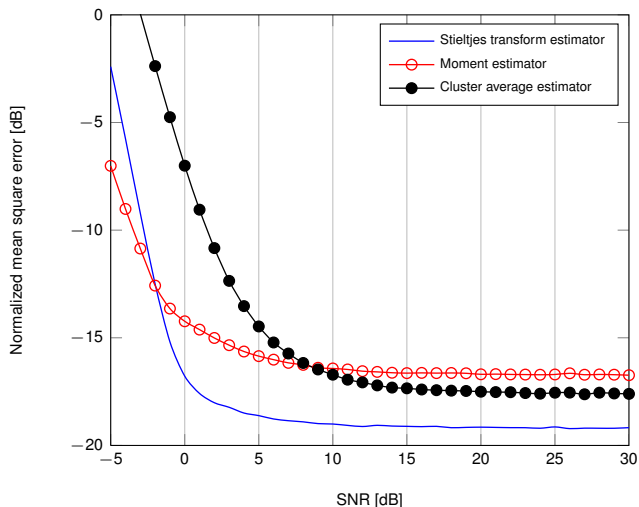


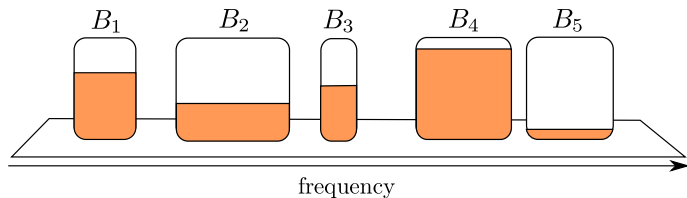
Figure: Normalized mean square error of largest estimated power \hat{P}_3 , $P_1 = 1/16$, $P_2 = 1/4$, $P_3 = 1$, $n_1 = n_2 = n_3 = 4$, $N = 24$, $M = 128$. Comparison between classical, moment and Stieltjes transform approaches.

Outline

- 1 Cognitive Radios
- 2 Exploration: Spectrum Hole Detection
- 3 Exploration: User Detection and Power Inference
- 4 Exploitation: Optimal Ergodic Rate**
- 5 Perspectives and Conclusion

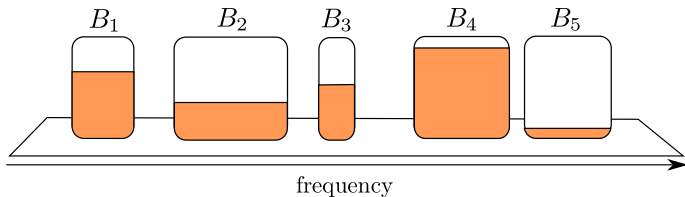
Problem statement

- We assume, from the exploration phase, that power Q_f can be transmitted in bandwidth B_f .



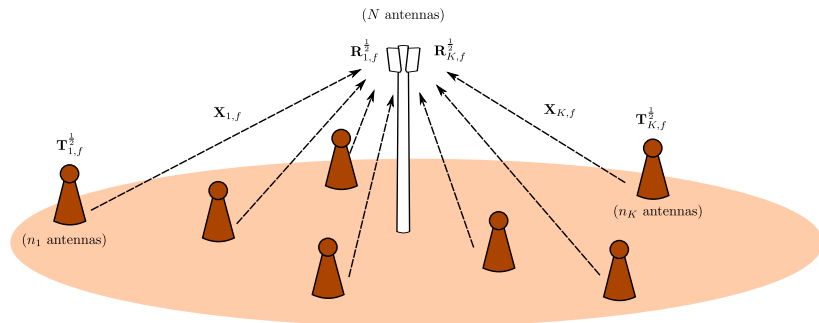
Problem statement

- We assume, from the exploration phase, that **power Q_f can be transmitted in bandwidth B_f .**
- The exploitation phase consists in **optimally using these power resources.**



Problem statement

- We assume, from the exploration phase, that **power Q_f can be transmitted in bandwidth B_f** .
- The exploitation phase consists in **optimally using these power resources**.
- We consider the uplink scenario of
 - an N -antenna base station;
 - K users equipped with n_1, \dots, n_K antennas;
 - Kronecker channels at all pairs $\mathbf{H}_{k,f} = \mathbf{R}_{k,f}^{\frac{1}{2}} \mathbf{X}_{k,f} \mathbf{T}_{k,f}^{\frac{1}{2}} \in \mathbb{C}^{N \times n_k}$ at frequency B_f ;
 - colored noise with covariance Σ .



Maximizing ergodic sum rate

- Due to mobility, we wish to optimize the **uplink ergodic sum rate** (per antenna),

$$\mathcal{I}(\mathbf{P}_{1,1}, \dots, \mathbf{P}_{K,F}) \triangleq \frac{1}{N} \sum_{f=1}^F \frac{|B_f|}{\sum_{f'} |B_{f'}|} \mathbb{E} \left[\log \det \left(\mathbf{I}_N + \sum_{k=1}^K \boldsymbol{\Sigma}_f^{-\frac{1}{2}} \mathbf{H}_{k,f} \mathbf{P}_{k,f} \mathbf{H}_{k,f}^H \boldsymbol{\Sigma}_f^{-\frac{1}{2}} \right) \right].$$

Maximizing ergodic sum rate

- Due to mobility, we wish to optimize the **uplink ergodic sum rate** (per antenna),

$$\mathcal{I}(\mathbf{P}_{1,1}, \dots, \mathbf{P}_{K,F}) \triangleq \frac{1}{N} \sum_{f=1}^F \frac{|B_f|}{\sum_{f'} |B_{f'}|} \mathbb{E} \left[\log \det \left(\mathbf{I}_N + \sum_{k=1}^K \Sigma_f^{-\frac{1}{2}} \mathbf{H}_{k,f} \mathbf{P}_{k,f} \mathbf{H}_{k,f}^H \Sigma_f^{-\frac{1}{2}} \right) \right].$$

- Determine the **sum rate maximizing precoders** $\mathbf{P}_{k,f}^*$

$$(\mathbf{P}_{1,1}^*, \dots, \mathbf{P}_{K,F}^*) = \arg \max_{\substack{\{\mathbf{P}_{k,f}\} \\ \sum_{k=1}^K \text{tr} \mathbf{P}_{k,f} \leq Q_f}} \mathcal{I}(\mathbf{P}_{1,1}, \dots, \mathbf{P}_{K,F}).$$

Maximizing ergodic sum rate

- Due to mobility, we wish to optimize the **uplink ergodic sum rate** (per antenna),

$$\mathcal{I}(\mathbf{P}_{1,1}, \dots, \mathbf{P}_{K,F}) \triangleq \frac{1}{N} \sum_{f=1}^F \frac{|B_f|}{\sum_{f'} |B_{f'}|} \mathbb{E} \left[\log \det \left(\mathbf{I}_N + \sum_{k=1}^K \Sigma_f^{-\frac{1}{2}} \mathbf{H}_{k,f} \mathbf{P}_{k,f} \mathbf{H}_{k,f}^H \Sigma_f^{-\frac{1}{2}} \right) \right].$$

- Determine the **sum rate maximizing precoders** $\mathbf{P}_{k,f}^*$

$$(\mathbf{P}_{1,1}^*, \dots, \mathbf{P}_{K,F}^*) = \arg \max_{\substack{\{\mathbf{P}_{k,f}\} \\ \sum_{k=1}^K \text{tr} \mathbf{P}_{k,f} \leq Q_f}} \mathcal{I}(\mathbf{P}_{1,1}, \dots, \mathbf{P}_{K,F}).$$

- Simplifying assumptions:

- Problem can be treated for each f independently. We then assume $F = 1$.
- Taking $\Sigma = \sigma^2 \mathbf{I}_N$ does not restrict generality.

Deterministic Equivalents of the Sum Rate

- The stochastic character of $\mathbf{H}_{k,f}$ makes things difficult.
- We instead find a **deterministic approximation** $\mathcal{I}^\circ(\mathbf{P}_1, \dots, \mathbf{P}_K)$ for $\mathcal{I}(\mathbf{P}_1, \dots, \mathbf{P}_K)$ such that

$$\mathcal{I}^\circ(\mathbf{P}_1, \dots, \mathbf{P}_K) - \mathcal{I}(\mathbf{P}_1, \dots, \mathbf{P}_K) \rightarrow 0$$

as $N, n_1, \dots, n_K \rightarrow \infty$, and denote

$$(\mathbf{P}_1^\circ, \dots, \mathbf{P}_K^\circ) = \arg \max_{\{\mathbf{P}_k\}} \mathcal{I}^\circ(\mathbf{P}_1, \dots, \mathbf{P}_K).$$

Deterministic Equivalents of the Sum Rate

- The stochastic character of $\mathbf{H}_{k,f}$ makes things difficult.
- We instead find a **deterministic approximation** $\mathcal{I}^\circ(\mathbf{P}_1, \dots, \mathbf{P}_K)$ for $\mathcal{I}(\mathbf{P}_1, \dots, \mathbf{P}_K)$ such that

$$\mathcal{I}^\circ(\mathbf{P}_1, \dots, \mathbf{P}_K) - \mathcal{I}(\mathbf{P}_1, \dots, \mathbf{P}_K) \rightarrow 0$$

as $N, n_1, \dots, n_K \rightarrow \infty$, and denote

$$(\mathbf{P}_1^\circ, \dots, \mathbf{P}_K^\circ) = \arg \max_{\{\mathbf{P}_k\}} \mathcal{I}^\circ(\mathbf{P}_1, \dots, \mathbf{P}_K).$$

- We can show

$$\mathcal{I}(\mathbf{P}_1^*, \dots, \mathbf{P}_K^*) - \mathcal{I}(\mathbf{P}_1^\circ, \dots, \mathbf{P}_K^\circ) \rightarrow 0.$$

Deterministic Equivalents: Strategy

- With $\mathbf{B}_N \triangleq \sum_{k=1}^K \mathbf{H}_k \mathbf{P}_k \mathbf{H}_k^H$ ($\mathbf{H}_k = \mathbf{R}_k^{\frac{1}{2}} \mathbf{X}_k \mathbf{T}_k^{\frac{1}{2}}$), notice that

$$\begin{aligned}
 \mathcal{I}(\mathbf{P}_1, \dots, \mathbf{P}_K) &= \mathbb{E} \left[\frac{1}{N} \log \det \left(\mathbf{I}_N + \frac{1}{\sigma^2} \mathbf{B}_N \right) \right] \\
 &= \mathbb{E} \left[\int_{\sigma^2}^{\infty} \left(\frac{1}{\omega} - \frac{1}{N} \operatorname{tr} (\mathbf{B}_N + \omega \mathbf{I}_N)^{-1} \right) d\omega \right] \\
 &= \mathbb{E} \left[\int_{\sigma^2}^{\infty} \left(\frac{1}{\omega} - m_{\mathbf{B}_N}(-\omega) \right) d\omega \right].
 \end{aligned}$$

Deterministic Equivalents: Strategy

- With $\mathbf{B}_N \triangleq \sum_{k=1}^K \mathbf{H}_k \mathbf{P}_k \mathbf{H}_k^H$ ($\mathbf{H}_k = \mathbf{R}_k^{\frac{1}{2}} \mathbf{X}_k \mathbf{T}_k^{\frac{1}{2}}$), notice that

$$\begin{aligned}
 \mathcal{I}(\mathbf{P}_1, \dots, \mathbf{P}_K) &= \mathbb{E} \left[\frac{1}{N} \log \det \left(\mathbf{I}_N + \frac{1}{\sigma^2} \mathbf{B}_N \right) \right] \\
 &= \mathbb{E} \left[\int_{\sigma^2}^{\infty} \left(\frac{1}{\omega} - \frac{1}{N} \operatorname{tr} (\mathbf{B}_N + \omega \mathbf{I}_N)^{-1} \right) d\omega \right] \\
 &= \mathbb{E} \left[\int_{\sigma^2}^{\infty} \left(\frac{1}{\omega} - m_{\mathbf{B}_N}(-\omega) \right) d\omega \right].
 \end{aligned}$$

- It suffices to find a **deterministic equivalent** for $m_{\mathbf{B}_N}(z)$.

Final Results

R. Couillet, M. Debbah and J. W. Silverstein, "A Deterministic Equivalent for the Analysis of Correlated MIMO Multiple Access Channels", *to appear in IEEE Transactions on Information Theory*, arXiv Preprint 0906.3667.

Theorem (Deterministic equivalent of the Stieltjes transform)

Under some mild conditions on the \mathbf{R}_k and \mathbf{T}_k matrices, as $N, n_1, \dots, n_K \rightarrow \infty$

$$m_{\mathbf{B}_N}(z) - m_N(z) \xrightarrow{\text{a.s.}} 0$$

where

$$m_N(z) = \frac{1}{N} \operatorname{tr} \left(-z \left[\sum_{k=1}^K \bar{e}_k(z) \mathbf{R}_k + \mathbf{I}_N \right] \right)^{-1}$$

and $\{\bar{e}_i(z)\}$, $i \in \{1, \dots, K\}$, form the unique solution to

$$e_i(z) = \frac{1}{n_i} \operatorname{tr} \mathbf{R}_i \left(-z \left[\sum_{k=1}^K \bar{e}_k(z) \mathbf{R}_k + \mathbf{I}_N \right] \right)^{-1}$$

$$\bar{e}_i(z) = \frac{1}{n_i} \operatorname{tr} \mathbf{T}_i^{\frac{1}{2}} \mathbf{P}_i \mathbf{T}_i^{\frac{1}{2}} \left(-z \left[e_i(z) \mathbf{T}_i^{\frac{1}{2}} \mathbf{P}_i \mathbf{T}_i^{\frac{1}{2}} + \mathbf{I}_{n_i} \right] \right)^{-1}$$

Final results (2)

R. Couillet, M. Debbah and J. W. Silverstein, "A Deterministic Equivalent for the Analysis of Correlated MIMO Multiple Access Channels", *to appear in IEEE Transactions on Information Theory*, arXiv Preprint 0906.3667.

Theorem (Deterministic equivalents of the sum rate)

Under similar conditions, with $\mathcal{I}(\mathbf{P}_1, \dots, \mathbf{P}_K) = \frac{1}{N} \mathbb{E} \log \det \left(\mathbf{I}_N + \frac{1}{\sigma^2} \mathbf{B}_N \right)$ and $z = -\sigma^2$,

$$\mathcal{I}(\mathbf{P}_1, \dots, \mathbf{P}_K) - \mathcal{I}^\circ(\mathbf{P}_1, \dots, \mathbf{P}_K) \rightarrow 0$$

where

$$\mathcal{I}^\circ(\mathbf{P}_1, \dots, \mathbf{P}_K) = \frac{1}{N} \log \left| \mathbf{I}_N + \sum_{k=1}^K \bar{\mathbf{e}}_k \mathbf{R}_k \right| + \sum_{k=1}^K \frac{1}{N} \log \left| \mathbf{I}_{n_k} + \mathbf{e}_k \mathbf{T}_k^{\frac{1}{2}} \mathbf{P}_k \mathbf{T}_k^{\frac{1}{2}} \right| - \sigma^2 \sum_{k=1}^K \frac{n_k}{N} \bar{\mathbf{e}}_k \mathbf{e}_k.$$

Final results (2)

R. Couillet, M. Debbah and J. W. Silverstein, "A Deterministic Equivalent for the Analysis of Correlated MIMO Multiple Access Channels", *to appear in IEEE Transactions on Information Theory*, arXiv Preprint 0906.3667.

Theorem (Deterministic equivalents of the sum rate)

Under similar conditions, with $\mathcal{I}(\mathbf{P}_1, \dots, \mathbf{P}_K) = \frac{1}{N} \mathbb{E} \log \det \left(\mathbf{I}_N + \frac{1}{\sigma^2} \mathbf{B}_N \right)$ and $z = -\sigma^2$,

$$\mathcal{I}(\mathbf{P}_1, \dots, \mathbf{P}_K) - \mathcal{I}^\circ(\mathbf{P}_1, \dots, \mathbf{P}_K) \rightarrow 0$$

where

$$\mathcal{I}^\circ(\mathbf{P}_1, \dots, \mathbf{P}_K) = \frac{1}{N} \log \left| \mathbf{I}_N + \sum_{k=1}^K \bar{\mathbf{e}}_k \mathbf{R}_k \right| + \sum_{k=1}^K \frac{1}{N} \log \left| \mathbf{I}_{n_k} + \mathbf{e}_k \mathbf{T}_k^{\frac{1}{2}} \mathbf{P}_k \mathbf{T}_k^{\frac{1}{2}} \right| - \sigma^2 \sum_{k=1}^K \frac{n_k}{N} \bar{\mathbf{e}}_k \mathbf{e}_k.$$

- It remains to maximize $\mathcal{I}^\circ(\mathbf{P}_1, \dots, \mathbf{P}_K)$ over $\mathbf{P}_1, \dots, \mathbf{P}_K$ with $\sum_k \text{tr} \mathbf{P}_k \leq Q$.

Final results (2)

R. Couillet, M. Debbah and J. W. Silverstein, "A Deterministic Equivalent for the Analysis of Correlated MIMO Multiple Access Channels", *to appear in IEEE Transactions on Information Theory*, arXiv Preprint 0906.3667.

Theorem (Deterministic equivalents of the sum rate)

Under similar conditions, with $\mathcal{I}(\mathbf{P}_1, \dots, \mathbf{P}_K) = \frac{1}{N} \mathbb{E} \log \det \left(\mathbf{I}_N + \frac{1}{\sigma^2} \mathbf{B}_N \right)$ and $z = -\sigma^2$,

$$\mathcal{I}(\mathbf{P}_1, \dots, \mathbf{P}_K) - \mathcal{I}^\circ(\mathbf{P}_1, \dots, \mathbf{P}_K) \rightarrow 0$$

where

$$\mathcal{I}^\circ(\mathbf{P}_1, \dots, \mathbf{P}_K) = \frac{1}{N} \log \left| \mathbf{I}_N + \sum_{k=1}^K \bar{\mathbf{e}}_k \mathbf{R}_k \right| + \sum_{k=1}^K \frac{1}{N} \log \left| n_k + \mathbf{e}_k \mathbf{T}_k^{\frac{1}{2}} \mathbf{P}_k \mathbf{T}_k^{\frac{1}{2}} \right| - \sigma^2 \sum_{k=1}^K \frac{n_k}{N} \bar{\mathbf{e}}_k \mathbf{e}_k.$$

- It remains to maximize $\mathcal{I}^\circ(\mathbf{P}_1, \dots, \mathbf{P}_K)$ over $\mathbf{P}_1, \dots, \mathbf{P}_K$ with $\sum_k \text{tr} \mathbf{P}_k \leq Q$.
- This is obtained by **iterative waterfilling**
 - \mathbf{P}_k and \mathbf{T}_k have the same eigenspaces
 - with $\text{eig}(\mathbf{P}_k^\circ) = (p_{k,1}^\circ, \dots, p_{k,n_k}^\circ)$

$$p_{k,i}^\circ = \left(\mu - \frac{1}{\mathbf{e}_k^\circ \mathbf{t}_{k,i}} \right)^+, \quad \mathbf{e}_k^\circ = \mathbf{e}_k(\mathbf{P}_1^\circ, \dots, \mathbf{P}_K^\circ), \quad \mu \text{ set to satisfy } \sum_k \text{tr} \mathbf{P}_k = Q.$$

Simulation results

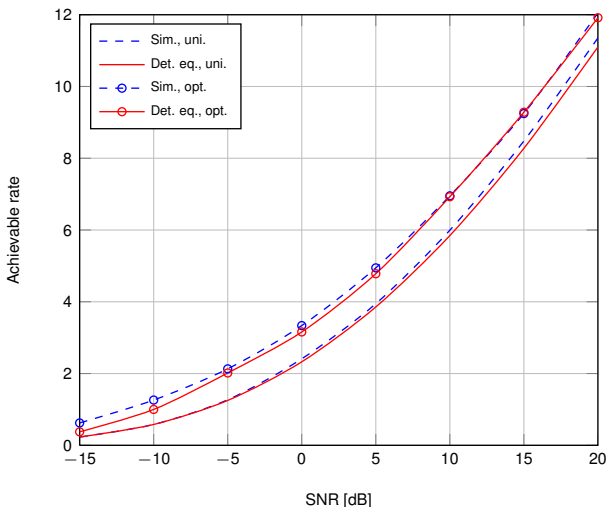


Figure: Ergodic MAC sum rate for an $N = 4$ antenna receiver and $K = 4$ single-antenna transmitters under sum power constraint. Every user transmit signal has different correlation patterns at the receiver, and different path losses. Deterministic equivalents (det. eq.) against simulation (sim.), with uniform (uni.) or optimal (opt.) power allocation.

Outline

- 1 Cognitive Radios
- 2 Exploration: Spectrum Hole Detection
- 3 Exploration: User Detection and Power Inference
- 4 Exploitation: Optimal Ergodic Rate
- 5 **Perspectives and Conclusion**

The road ahead

- **signal sensing:**

- optimal hypothesis tests require **symmetry**, are often **computationally prohibitive**;
- situations with more side information demand simpler tests, based on eigen-structure;
- more realistic scenarios with cooperation will demand a further improvement of such tests.

The road ahead

● **signal sensing:**

- optimal hypothesis tests require **symmetry**, are often **computationally prohibitive**;
- situations with more side information demand simpler tests, based on eigen-structure;
- more realistic scenarios with cooperation will demand a further improvement of such tests.

● **statistical inference:**

- many methods have been proposed recently (moments, direct inversion, Stieltjes transform . . .)
 - Stieltjes transform approach seems the most powerful;
 - Stieltjes transform approach **suffers when separability is lost**;
 - **estimating number of sources remains**.
- **test performance** must be better evaluated;
- there is **room for extension** to more realistic/involved models.

The road ahead

● **signal sensing:**

- optimal hypothesis tests require **symmetry**, are often **computationally prohibitive**;
- situations with more side information demand simpler tests, based on eigen-structure;
- more realistic scenarios with cooperation will demand a further improvement of such tests.

● **statistical inference:**

- many methods have been proposed recently (moments, direct inversion, Stieltjes transform . . .)
 - Stieltjes transform approach seems the most powerful;
 - Stieltjes transform approach **suffers when separability is lost**;
 - **estimating number of sources remains**.
- **test performance** must be better evaluated;
- there is **room for extension** to more realistic/involved models.

● **overlaid communications:**

- **capacity evaluation in multi-dimensional networks** is progressing fast;
- optimal feedback and cooperation needs to be developed with similar random matrix tools.

The road ahead

● **signal sensing:**

- optimal hypothesis tests require **symmetry**, are often **computationally prohibitive**;
- situations with more side information demand simpler tests, based on eigen-structure;
- more realistic scenarios with cooperation will demand a further improvement of such tests.

● **statistical inference:**

- many methods have been proposed recently (moments, direct inversion, Stieltjes transform . . .)
 - Stieltjes transform approach seems the most powerful;
 - Stieltjes transform approach **suffers when separability is lost**;
 - **estimating number of sources remains**.
- **test performance** must be better evaluated;
- there is **room for extension** to more realistic/involved models.

● **overlaid communications:**

- **capacity evaluation in multi-dimensional networks** is progressing fast;
- optimal feedback and cooperation needs to be developed with similar random matrix tools.

Future cognitive radio communications with

- open-access communications,
- secondary networks coordination,

will demand

- a characterization of **what information to share**,
- a merger between **random matrix theory and game theory**,
- a stronger effort on **graph-oriented random matrix theory**.

● **Publications in Journals: (7 accepted / 2 submitted)**

- **R. Couillet**, J. Hoydis, M. Debbah, “A deterministic equivalent approach to the performance analysis of isometric random precoded systems,” *submitted to IEEE Transactions on Information Theory*, arXiv Preprint XXXX.XXXX.
- S. Wagner, **R. Couillet**, M. Debbah, D. T. M. Slock, “Large System Analysis of Linear Precoding in MISO Broadcast Channels with Limited Feedback”, *submitted to IEEE Transactions on Information Theory*, arXiv Preprint 0906.3682.
- **R. Couillet**, J. W. Silverstein, Z. Bai, M. Debbah, “Eigen-Inference for Energy Estimation of Multiple Sources”, *to appear in IEEE Transactions on Information Theory*, arXiv Preprint 1001.3934.
- **R. Couillet**, M. Debbah, J. W. Silverstein, “A Deterministic Equivalent for the Analysis of Correlated MIMO Multiple Access Channels”, *to appear in IEEE Transactions on Information Theory*, arXiv Preprint 0906.3667.
- **R. Couillet**, M. Debbah, “A Bayesian Framework for Collaborative Multi-Source Signal Sensing”, *IEEE Transactions on Signal Processing*, vol. 58, no. 10, pp. 5186-5195, 2010, arXiv Preprint 0811.0764.
- **R. Couillet**, A. Ancora, M. Debbah, “Bayesian Foundations of Channel Estimation for Cognitive Radios”, *Advances in Electronics and Telecommunications*, vol. 1, no. 1, pp. 41-49, 2010.
- **R. Couillet**, M. Debbah, “Le téléphone du futur : plus intelligent pour une exploitation optimale des fréquences” *Revue de l'Electricité et de l'Electronique*, no. 6, pp. 71-83, 2010.
- **R. Couillet**, M. Debbah, “Mathematical foundations of cognitive radios”, *Journal of Telecommunications and Information Technologies*, no. 4, 2009.
- **R. Couillet**, M. Debbah, “Outage performance of flexible OFDM schemes in packet-switched transmissions”, *Eurasip Journal on Advances on Signal Processing*, Volume 2009, Article ID 698417, 2009.

● Publications in International Conferences: (16 accepted / 3 submitted)

- A. Kammoun, **R. Couillet**, J. Najim, M. Debbah, "A G-estimator for rate adaption in cognitive radios," *submitted to IEEE Dynamic Spectrum Access Networks*, Aachen, Germany, 2011.
- J. Yao, **R. Couillet**, J. Najim, E. Moulines, M. Debbah, "CLT for eigen-inference methods in cognitive radios," *submitted to IEEE International Conference on Acoustics, Speech and Signal Processing*, Prague, Czech Republic, 2011.
- J. Hoydis, **R. Couillet**, M. Debbah, "Deterministic Equivalents for the Performance Analysis of Isometric Random Precoded Systems," *submitted to IEEE International Conference on Communications*, Kyoto, Japan, 2011.
- J. Hoydis, J. Najim, **R. Couillet**, M. Debbah, "Fluctuations of the Mutual Information in Large Distributed Antenna Systems with Colored Noise," Forty-Eighth Annual Allerton Conference on Communication, Control, and Computing, Allerton, IL, USA, 2010.
- **R. Couillet**, H. V. Poor, M. Debbah, "Self-organized spectrum sharing in large MIMO multiple-access channels," IEEE International Symposium on Information Theory, Austin TX, USA, 2010.
- **R. Couillet**, J. W. Silverstein, M. Debbah, "Eigen-inference for multi-source power estimation," IEEE International Symposium on Information Theory, Austin TX, USA, 2010.
- S. Wagner, **R. Couillet**, D. T. M. Slock, M. Debbah, "Optimal Training in Large TDD Multi-user Downlink Systems under Zero-forcing and Regularized Zero-forcing Precoding," *submitted to IEEE Global Communication Conference*, Miami, 2010.
- S. Wagner, **R. Couillet**, D. T. M. Slock, M. Debbah, "Large System Analysis of Zero-Forcing Precoding in MISO Broadcast Channels with Limited Feedback" IEEE International Workshop on Signal Processing Advances for Wireless Communications, Marrakech, Morocco, 2010.
- **R. Couillet**, M. Debbah, "Information theoretic approach to synchronization: the OFDM carrier frequency offset example", Advanced International Conference on Telecommunications, Barcelona, Spain, 2010.
- **R. Couillet**, M. Debbah, "Uplink capacity of self-organizing clustered orthogonal CDMA networks in flat fading channels" IEEE Information Theory Workshop Fall'09, Taormina, Sicily, 2009.
- **R. Couillet**, M. Debbah, J. W. Silverstein, "Asymptotic Capacity of Multi-User MIMO Communications" IEEE Information Theory Workshop Fall'09, Taormina, Sicily, 2009.
- **R. Couillet**, M. Debbah, J. W. Silverstein, "Rate region of correlated MIMO multiple access channel and broadcast channel" IEEE Workshop on Statistical Signal Processing, Cardiff, Wales, UK, 2009.
- **R. Couillet**, M. Debbah, "Mathematical foundations of cognitive radios" U.R.S.I.'09, Warsaw, Poland, 2009.
- **R. Couillet**, M. Debbah, "A maximum entropy approach to OFDM channel estimation", IEEE International Workshop on Signal Processing Advances for Wireless Communications, Perugia, Italy, 2009.
- **R. Couillet**, M. Debbah, "Bayesian inference for multiple antenna cognitive receivers", IEEE Wireless Communications & Networking Conference, Budapest, Hungary, 2009.
- **R. Couillet**, M. Debbah, "Flexible OFDM schemes for bursty transmissions", IEEE Wireless Communications & Networking Conference, Budapest, Hungary, 2009.
- **R. Couillet**, S. Wagner, M. Debbah, "Asymptotic Analysis of Correlated Multi-Antenna Broadcast Channels", IEEE Wireless Communications & Networking Conference, Budapest, Hungary, 2009.
- **R. Couillet**, S. Wagner, M. Debbah, A. Silva, "The Space Frontier: Physical Limits of Multiple Antenna Information Transfer", ValueTools, Inter-Perf Workshop, Athens, Greece, 2008. **BEST STUDENT PAPER AWARD**
- **R. Couillet**, M. Debbah, "Free deconvolution for OFDM multicell SNR detection", IEEE Personal, Indoor and Mobile Radio Communications Symposium, Cognitive Radio Workshop, Cannes, France, 2008.

- **Patents: (4 owned by ST-Ericsson)**

- **R. Couillet**, M. Debbah, *Application no. 08368028.0* “Process and apparatus for performing initial carrier frequency offset in an OFDM communication system”
- **R. Couillet**, M. Debbah, *Application no. 08368023.1* “Method for short-time OFDM transmission and apparatus for performing flexible OFDM modulation”
- **R. Couillet**, S. Wagner, *Application no. 09368025.4* “Precoding process for a transmitter of a MU-MIMO communication system”
- **R. Couillet**, *Application no. 09368030.4* “Process for estimating the channel in an OFDM communication system, and receiver for doing the same”

- **Books and book chapters: (1 book / 2 book chapters)**

- **Random Matrix Methods for Wireless Communications** [book]

Theoretical random matrix tools (finite dimensional analysis, limiting spectral laws, free probability, deterministic equivalents, statistical inference) and applications to wireless communications (SU-MIMO, MU-MIMO, CDMA, detection, estimation, channel modelling).

- *Authors:* **R. Couillet** and M. Debbah
- *Publisher:* Cambridge University Press
- *Year:* 2011 (to appear)

- **Mathematical Foundations for Signal Processing, Communications and Networking** [book chapter XX]

Chapter “Random matrix theory” on reminders of random matrix theory and especially statistical inference methods.

- *Chapter authors:* **R. Couillet** and M. Debbah
- *Publisher:* CRC Press, Taylor & Francis Group
- *Year:* 2011 (to appear)

- **Orthogonal Frequency Division Multiple Access Fundamentals and Applications** [book chapter 13]

Chapter “Fundamentals of OFDMA Synchronization” on theoretical considerations and application tools for time offset and frequency offset regulation in OFDM and OFDMA systems.

- *Chapter authors:* **R. Couillet** and M. Debbah
- *Publisher:* Auerbach Publications, CRC Press, Taylor & Francis Group
- *ISBN:* 978-1-4200-8824-3
- *Year:* 2010

Thank you for your attention.